# **Glitch Rejection Properties of a Simple Coherent Network Algorithm for Detecting Unmodeled GW Bursts**



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## Introduction

In GWB detection we face a twofold difficulty. First the sought waveforms are unmodeled; second, transient instrumental disturbances (glitches) make the noise non-Gaussian and poorly statistically modeled. Coherent network algorithms exploit network redundancy to estimate the waveforms from the data; suitable versions of the likelihood ratio are accordingly obtained, usually based on the Gaussian background noise assumption, and adhoc arguments are added aimed at assessing the ability to discriminate spurious glitches from true GWBs. Here we propose a general promising model for the impulsive IFO noise component along the lines laid out by D. Middleton. We use this model to assess the performance of the cross-correlation based coherent detection algorithm proposed by Rakhmanov and Klimenko (RK), for the case of triggered search, highlighting the effect of glitches in terms of ROC degradation. The results are obtained in a readable analytic form.

#### Middleton's model for glitches

Experimental evidence suggests that instrumental noise glitches can be modelled as atoms (waveforms with almost-compact support) in the time-frequency (TF) plane, characterized in terms of energy content, occurrence time  $t_0$ , center frequency  $f_0$ , effective duration  $\sigma_t$  and bandwidth  $\sigma_r$ 

The impulsive noise component, g(t), can be modelled as a random process consisting of a linear superposition of atoms, viz.

$$g(t) = \sum_{k=1}^{K[T]} \psi(t - t_0^{(k)}; \underline{a}^{(k)}), \quad t \in \Theta$$

- chosen representation atom;
- $\begin{array}{c}\psi(\cdot)\\t_0^{(k)}\\\underline{a}^{(k)}\end{array}$ glitch firing times (random);
- atoms' shape vectors (random ); K[T]# of glitches in (T seconds wide) analysis
- window  $\theta$  (random).

The random parameters  $a^{(k)}$  are determined inde*pendently* at each glitch instance (i.e. for each k); their distribution must be deduced from available data: K[T] follows a Poisson distribution [Hurwitz and Kac, Ann. Math. Stat. 15 (1944) 173], i.e.

prob {
$$K[T] = K$$
} =  $\overbrace{K!}{K!} = \overbrace{K!}{K!}$ 

The characteristic function of g(t) can be computed exactly up to any order [D. Middleton, J. Appl. Phys. 22 (1951) 1143]. To first order

 $F_{g}(\xi,t) = \exp\left[\overline{N}\left(E\left[e^{i\xi\psi(t-t_{0};\underline{a})}\right]-1\right)\right]$ 

yielding for the first two moments

$$\mu_g^{(1)} = E[g(t)] = \overline{N}E[\psi(t-t_0;\underline{a})]_{t_0,\underline{a}}$$

$$\mu_g^{(2)} = E\left[g(t)^2\right] = \overline{N}^2 E^2\left[\psi\left(t-t_0;\underline{a}\right)\right]_{t_0,a} + \overline{N}E\left[\psi^2\left(t-t_0;\underline{a}\right)\right]_{t_0,a}$$

Moments are time independent, such being by assumption the prior distributions of  $t_0^{(k)}$ ,  $a^{(k)}$  in  $\Theta$ . Here we adopt the (real) Sine-Gaussian (SG) atom, in view of its structural simplicity and minimum TF spread property ( $\sigma_t \sigma_f = 1/4\pi$ )

$$\psi(t-t_0; g_0, f_0, \sigma_t) = g_0 \sin \left[ 2\pi f_0(t-t_0) \right] e^{-(t-t_0)^2 / \sigma_t^2}$$

for which



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## The RK Correlation Algorithm

As a preparation to the more general (and complicated) case where one works in terms of the network likelihood ratio, we consider the simplest detection algo for unmodelled bursts devised by Rakhmanov and Klimenko [CQG 22 (2005) S1311], exploiting network redundancy to first construct noisy templates for detector i

 $W_i(t) = -A_i V_i(t) - A_\ell V_\ell(t)$ 

or/from properly time-shifted 
$$(\tau_i)$$
 IFO outputs

$$V_i(t) = S_i(t) + n_i(t)$$

where  $A_i = F_i^+ F_\ell^- - F_i^* F_\ell^+$ ,  $i \neq j \neq \ell$  and then to compute the overlaps (correlators):

$$C_i = \langle V_i, W_i \rangle \approx f_s^{-1} \sum_{i=1}^{N_i} V_{im} W_i$$

These are the best measure of overlap between  $V_i$ and  $W_i$  for Gaussian background noise, and still asymptotically optimal in the weak signal limit, for non-Gaussian noise, up to a suitable non-linear preprocessing of data ( $V_i$ ), to be discussed elsewhere. Pattern factors  $F_i^{+,\times}$  and propagation delays  $\tau_i$  depend on the direction of arrival (DOA)  $\Omega_{\rm s}$  . DOA and time of occurrence assumed known (triggered search).

Irrespective of the underlying noise properties, the RK correlators are Gaussian distributed (CLT) and thus completely characterized by their 1st and 2nd order moments

Moments under 
$$H_I$$
 hypothesis ...  
 $N_t$ :# of samples in  $\Theta$ 

$$\left(\sigma_{1}^{(i)}\right)^{2} = Var\left[C_{i}|H_{1}\right] \approx f_{s}^{-2}A_{i}^{2}\left[\left(\sigma_{i}^{2} + \tilde{\sigma}_{i}^{2}\right)\mathbf{S}_{i} \cdot \mathbf{S}_{i}^{T} + \sigma_{i}^{2}\tilde{\sigma}_{i}^{2}N_{s}\right]$$

where 
$$\sigma$$
 is the noise variance in the detectors, and  $\tilde{\sigma}_i^2 = (A_i^2 + A_t^2) A_t^{-2} \sigma_t^2$  is the noise variance in the tem-

plate.  $\begin{array}{l} \delta_{S}: \text{SNR of } V_{i} \\ \delta_{k}: \text{SNR of GW signal} \\ |F_{i}|: \text{ antenna response at } \Omega_{s} \end{array}$ ... corresponding deflection  $A_i$ 

$$I^{(i)} = \delta_{S}^{(i)} \frac{A_{i}}{\left[A_{i}^{2} + A_{j}^{2} + A_{\ell}^{2} + N_{s}\left(A_{j}^{2} + A_{\ell}^{2}\right)\left(|F_{i}|\delta_{k}\right)^{-2}\right]^{1/2}}, \quad \delta_{S}^{(i)} = |F_{i}|\delta_{k}$$

GWB and glitches occurrence in  $\Theta$  ruled out by hyp. Moments under  $H_{a}$  hypothesis  $2A_{A}$  $\Xi(\Omega_{,}) =$  $\overline{A_i^2 + A_\ell^2}$  $\mu_0^{(i)} = E \left[ C_i | H_0 \right] = -f_s^{-1} N_s \left( A_i + A_\ell \right) \left( \mu_g^{(1)} \right)^2$ 

$$\left(\sigma_{0}^{(i)}\right)^{2} = Var\left[C_{i} \mid H_{0}\right] = \left[\sigma_{AHGA}^{(i)}\right)^{2} \left\{1 + 2\left[\mu_{g}^{(2)} + \Xi\left(\Omega_{s} \left(\frac{\mu_{g}^{(1)}}{\sigma^{2}}\right)^{2}\right] \right] \right\}$$
  
AWGN term

We consider the network: {LHO (H), LLO (L), Virgo (V)}. From the Gaussianity of  $C_i$ , the false dismissal probability  $\beta$  is related to the false alarm probability  $\alpha$  as follows (Receiver Operating Characteristics, ROC):

$$\beta^{(i)} = 1 - \operatorname{erfc}\left[\frac{\sigma_0^{(i)}}{\sigma_1^{(i)}} \operatorname{erfc}^{-1}(\alpha) + \frac{\mu_0^{(i)}}{\sigma_1^{(i)}} - \frac{\mu_1^{(i)}}{\sigma_1^{(i)}}\right]$$

## **Simulations**

Figs 1,2 illustrate the performance of the RK detector compared to the ideal matched filter (MF). A degradation by a factor ~2 in terms of deflection at the optimum DOA and by a factor ~3 averaged over all sky is observed. Glitches are neglected in Figs 1,2. The effects of glitches are illustrated in Figs 3,4, for different firing rates (Fig 3) and SNR<sub>veto</sub> levels (Fig 4)

The prior distributions of  $f_0$  and  $\sigma_{\rm t}$  for SG atoms, used in drawing Figs 3,4 were obtained from 1 week of instrumental triggers [http://ldas-jobs.ligo-wa. caltech.edu/qonline]. The glitch SNR is assumed as uniformly distributed between zero and SNR, Performance deteriorates as either the glitch firing rate or/and the SNR<sub>veto</sub> level increases.



## Conclusions/hints for future work Main results:

- New characterization of the impulsive IFO noise component based on atomic representation for glitches and Middleton's model;
- Fully analytical model for RK performance under CFAR operation detector's (triggered search), including glitches obtained;
- RK detector robustness against instrumental glitches illustrated in terms of ROCs.

#### Directions for future work:

- Better glitch modeling (better atom dictionary and more accurate characterization of glitch parameters priors).
- Better detection statistics (e.g., directionoptimized combinations of the  $C_i$ );
- Implementation of "quasi-optimal" detector tailored to non-Gaussianity by nonlinear data conditioning before correlation (exploiting fully Middleton's model).