

# The $\mathcal{F}$ -statistic and its implementation in `ComputeFstatistic_v2`

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2018-07-24 15:59:07 +0200; commitID: c5de9d0(CLEAN)

LIGO-T0900149-v6

## Abstract

These notes represent a somewhat high-level documentation of `ComputeFstatistic_v2`, starting from a derivation and general discussion of the  $\mathcal{F}$ -statistic, down to expressions that very closely resemble what is actually implemented in the code.

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# 1 The signal $h(t)$ measured at the detector

## 1.1 General waveform

A plane gravitational wave  $h_{\mu\nu}$  propagating along the unit-vector  $-\hat{n}$  can be written in TT gauge as a purely spatial tensor  $\underline{h}$ , namely

$$\underline{h}(t, \vec{r}) = h_+(\tau) \underline{e}^+ + h_\times(\tau) \underline{e}^\times, \quad (1)$$

where  $\tau = t + \hat{n} \cdot \vec{r}/c$  and the polarization tensors  $\underline{e}^{\{+, \times\}}$  are defined as

$$\underline{e}^+ = \hat{u} \otimes \hat{u} - \hat{v} \otimes \hat{v}, \quad \text{and} \quad \underline{e}^\times = \hat{u} \otimes \hat{v} + \hat{v} \otimes \hat{u}, \quad (2)$$

in terms of unit vectors  $\hat{u}, \hat{v}$  that form an orthonormal basis  $\{\hat{u}, \hat{v}, -\hat{n}\}$  of

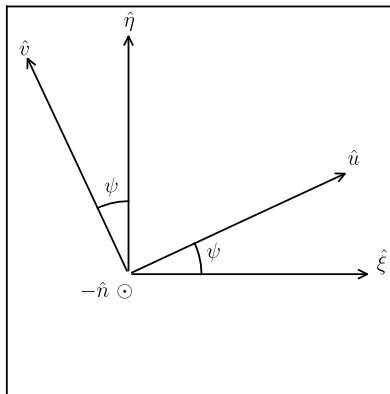


Figure 1: Illustration of different basis in the wave-plane, and definition of polarization angle  $\psi$  [Credit: John T Whelan]

the wave frame. The choice of basis  $\{\hat{u}, \hat{v}\}$  in the transversal wave plane is arbitrary, but one often chooses preferred directions given either by the source-geometry or by the principal polarization axis of elliptically polarized waves. It is therefore convenient to re-express this in a source-independent basis that only depends on the propagation direction  $-\hat{n}$  of the wave and the choice of an SSB-fixed reference frame  $\{\hat{x}, \hat{y}, \hat{z}\}$ . Such a frame is conventionally constructed using the unit basis vectors  $\hat{\xi} \equiv \hat{n} \times \hat{z}/|\hat{n} \times \hat{z}|$ ,  $\hat{\eta} \equiv \hat{\xi} \times \hat{n}$  and  $-\hat{n}$ . This definition is such that  $\hat{\xi}$  lies in the equatorial plane and  $\hat{\eta}$  points into the northern hemisphere. We now define the polarization angle  $\psi$  as the angle from  $\hat{\xi}$  to  $\hat{u}$ , measured counter-clockwise in the plane with  $-\hat{n}$  pointing at us<sup>1</sup>, i.e.  $\sin \psi = \hat{u} \cdot \hat{\eta}$ , see Fig. 1.

<sup>1</sup>This is what is meant by the phrase “counter-clockwise around  $-\hat{n}$ ” used in [8, 12]

This allows us to express the polarization basis  $\{\hat{u}, \hat{v}\}$  in terms of the basis  $\{\hat{\xi}, \hat{\eta}\}$  via a simple rotation by  $\psi$  around  $-\hat{n}$ , namely

$$\hat{u} = \hat{\xi} \cos \psi + \hat{\eta} \sin \psi, \quad (3)$$

$$\hat{v} = -\hat{\xi} \sin \psi + \hat{\eta} \cos \psi. \quad (4)$$

Introducing polarization-independent basis tensors in the wave-frame,

$$\underline{\varepsilon}^+ \equiv \hat{\xi} \otimes \hat{\xi} - \hat{\eta} \otimes \hat{\eta}, \quad (5)$$

$$\underline{\varepsilon}^\times \equiv \hat{\xi} \otimes \hat{\eta} + \hat{\eta} \otimes \hat{\xi}, \quad (6)$$

we can express the wave-basis  $\underline{e}^{\{+, \times\}}$  as

$$\underline{e}^+ = \cos 2\psi \underline{\varepsilon}^+ + \sin 2\psi \underline{\varepsilon}^\times \quad (7)$$

$$\underline{e}^\times = -\sin 2\psi \underline{\varepsilon}^+ + \cos 2\psi \underline{\varepsilon}^\times. \quad (8)$$

In the long-wavelength limit (LWL), where the arm length  $L$  of the detector satisfies  $L \ll \lambda/2\pi$  (where  $\lambda$  is the GW wavelength), the scalar response  $h^X(t)$  of a detector  $X$  to an incident GW tensor  $\underline{h}$  is expressible simply in terms of its detector tensor  $\underline{d}^X$ , namely

$$h^X(t) = \underline{d}^X(t) : \underline{h}(\tau^X) = d_{ij}^X h^{ij}(\tau^X), \quad (9)$$

where  $\tau^X(t) = t + \hat{n} \cdot \vec{r}^X(t)/c$  is (neglecting relativistic corrections) the arrival time of a wavefront at the SSB, which arrives at the detector  $X$  (at position  $\vec{r}^X$ ) at time  $t$ . This timing relation accounts for the Doppler effect due to the motion of the detector relative to the source. The LWL detector tensor for an interferometer with arms along  $\hat{l}_1$  and  $\hat{l}_2$  is simply given by

$$\underline{d} = \frac{1}{2} \left( \hat{l}_1 \otimes \hat{l}_1 - \hat{l}_2 \otimes \hat{l}_2 \right). \quad (10)$$

Using (1), we can write (9) in the form

$$h^X(t) = F_+^X(t) h_+(\tau^X) + F_\times^X(t) h_\times(\tau^X), \quad (11)$$

in terms of the so-called beam-pattern functions

$$F_+^X(t; \hat{n}, \psi) \equiv \underline{d}^X(t) : \underline{e}^+, \quad F_\times^X(t; \hat{n}, \psi) \equiv \underline{d}^X(t) : \underline{e}^\times. \quad (12)$$

Changing to the polarization-independent basis  $\underline{\varepsilon}^{+, \times}$  using (7), we find

$$F_+^X(t; \hat{n}, \psi) = a^X(t; \hat{n}) \cos 2\psi + b^X(t; \hat{n}) \sin 2\psi, \quad (13)$$

$$F_\times^X(t; \hat{n}, \psi) = b^X(t; \hat{n}) \cos 2\psi - a^X(t; \hat{n}) \sin 2\psi, \quad (14)$$

where the antenna-pattern functions  $a^X, b^X$  are defined as

$$a^X(t; \hat{n}) \equiv \underline{d}^X(t) : \underline{\varepsilon}^+(\hat{n}), \quad b^X(t; \hat{n}) \equiv \underline{d}^X(t) : \underline{\varepsilon}^\times(\hat{n}). \quad (15)$$

This formulation has the advantage of explicitly factoring out the polarization angle  $\psi$ .

As shown in [5][12], at the next order of approximation for longer armlength (the ‘‘rigid adiabatic approximation’’ RAA), the detector-tensor becomes frequency- and sky-location dependent, which formally can be absorbed into complex per-SFT antenna-pattern  $a^X$  and  $b^X$ , while the following  $\mathcal{F}$ -statistic expressions remain otherwise unchanged. The sky-position  $\hat{n}$  of the source is expressible in standard equatorial (or ecliptic) coordinates  $\alpha$  (right ascension), and  $\delta$  (declination) as

$$\hat{n} = (\cos \delta \cos \alpha, \cos \delta \sin \alpha, \sin \delta), \quad (16)$$

and by the above definitions, the corresponding polarization-independent wave-plane basis  $\hat{\xi}, \hat{\eta}$  is therefore expressible as

$$\begin{aligned} \hat{\xi} &= (\sin \alpha, -\cos \alpha, 0), \\ \hat{\eta} &= (-\cos \alpha \sin \delta, -\sin \alpha \sin \delta, \cos \delta). \end{aligned} \quad (17)$$

The contractions (15) are explicitly given by

$$\underline{d} : \underline{\varepsilon} = d_{11}\varepsilon_{11} + d_{22}\varepsilon_{22} + d_{33}\varepsilon_{33} + 2(d_{12}\varepsilon_{12} + d_{13}\varepsilon_{13} + d_{23}\varepsilon_{23}), \quad (18)$$

where  $\underline{\varepsilon}^{\{+, \times\}}$  are easily computed in SSB coordinates from (17), and the problem of computing  $a, b$  is therefore reduced to computing the detector tensor  $\underline{d}^X(t)$  as a function of time in this coordinate system.

See [11] for more detailed discussion and visualization of detector-tensor geometry.

## 1.2 Continuous-wave signals

The GW class of ‘‘continuous waves’’ is characterized by a signal model  $h_{+, \times}(\tau)$  (in the SSB) of the form

$$h_+(\tau) = A_+ \cos \Phi(\tau), \quad h_\times(\tau) = A_\times \sin \Phi(\tau). \quad (19)$$

Assuming a slowly varying intrinsic signal frequency  $2\pi f(\tau) \equiv d\Phi(\tau)/d\tau$ , the phase  $\Phi(\tau)$  can be expanded around the reference time  $\tau_{\text{ref}}$ , namely

$$\Phi(\tau) = \phi_0 + \phi(\Delta\tau), \quad \text{where} \quad (20)$$

$$\phi_0 \equiv \Phi(\tau_{\text{ref}}), \quad (21)$$

$$\phi(\Delta\tau) \equiv 2\pi \sum_{s=0} \frac{f^{(s)}(\tau_{\text{ref}})}{(s+1)!} (\Delta\tau)^{s+1}. \quad (22)$$

The detector-specific timing relation for isolated neutron stars contains relativistic corrections for the light-travel in the solar system. These corrections are taken into account in the numerical  $\mathcal{F}$ -statistic computation in `CFS_v2`, but for simplicity we give here only the first order Newtonian timing model,

$$\Delta\tau^X(t; \hat{n}) \equiv \tau^X - \tau_{\text{ref}} \approx t - \tau_{\text{ref}} + \frac{\vec{r}^X(t) \cdot \hat{n}}{c}, \quad (23)$$

where  $\tau^X$  is the arrival-time in the SSB of the GW phase reaching detector  $X$  at time  $t$ . The spin parameters  $f^{(s)}(\tau_{\text{ref}})$  are defined as

$$f^{(s)}(\tau_{\text{ref}}) \equiv \left. \frac{d^s f(\tau)}{d\tau^s} \right|_{\tau_{\text{ref}}}. \quad (24)$$

We denote the set of ‘‘Doppler parameters’’ affecting the time evolution of the phase  $\phi(\Delta\tau^X)$  as  $\lambda \equiv \{\hat{n}, f^{(s)}(\tau_{\text{ref}})\}$ . Combining (11), (13) (19), we find

$$h^X(t; \mathcal{A}, \lambda) = \sum_{\mu=1}^4 \mathcal{A}^\mu h_\mu^X(t; \lambda), \quad (25)$$

with the four amplitude parameters  $\mathcal{A}^\mu$  given by

$$\begin{aligned} \mathcal{A}^1 &= A_+ \cos \phi_0 \cos 2\psi - A_\times \sin \phi_0 \sin 2\psi, \\ \mathcal{A}^2 &= A_+ \cos \phi_0 \sin 2\psi + A_\times \sin \phi_0 \cos 2\psi, \\ \mathcal{A}^3 &= -A_+ \sin \phi_0 \cos 2\psi - A_\times \cos \phi_0 \sin 2\psi, \\ \mathcal{A}^4 &= -A_+ \sin \phi_0 \sin 2\psi + A_\times \cos \phi_0 \cos 2\psi, \end{aligned} \quad (26)$$

which is a re-parametrization of the (detector-independent) signal-parameters  $A_+, A_\times, \phi_0, \psi$ . The (detector-dependent) wave-components  $h_\mu^X(t; \lambda)$  are

$$\begin{aligned} h_1^X(t) &= a^X(t) \cos \phi(\Delta\tau^X), & h_2^X(t) &= b^X(t) \cos \phi(\Delta\tau^X), \\ h_3^X(t) &= a^X(t) \sin \phi(\Delta\tau^X), & h_4^X(t) &= b^X(t) \sin \phi(\Delta\tau^X). \end{aligned} \quad (27)$$

It is often useful to also consider the complex basis functions instead

$$\begin{aligned} h_a^X(t) &\equiv h_1^X - ih_3^X = a^X e^{-i\phi^X}, \\ h_b^X(t) &\equiv h_2^X - ih_4^X = b^X e^{-i\phi^X}. \end{aligned} \quad (28)$$

We see from (26) that there is some gauge-freedom in the amplitude-parameters  $\{A_+, A_\times, \psi, \phi_0\}$ , namely

$$\begin{aligned} \text{(i)} \quad &\psi \rightarrow \psi + \pi/2, \quad \phi_0 \rightarrow \phi_0 + \pi \\ \text{(ii)} \quad &\psi \rightarrow \psi + \pi/4, \quad \phi_0 \rightarrow \phi_0 - \pi/2, \quad A_+ \leftrightarrow A_\times \\ \text{(iii)} \quad &\phi_0 \rightarrow \phi_0 + \pi, \quad A_+ \rightarrow -A_+, \quad A_\times \rightarrow -A_\times \end{aligned} \quad (29)$$

Applying (i) twice, and taking account of the trivial gauge-freedom by  $2\pi$ , we also obtain the invariance  $\psi \rightarrow \psi + \pi$ .

In the case of a triaxial NS, the signal-amplitudes  $A_{+/\times}$  are expressible explicitly in terms of the wave-amplitude  $h_0$  and the inclination angle  $\iota$  with respect to the line-of-sight, namely

$$A_+ = \frac{1}{2}h_0 (1 + \cos^2 \iota), \quad A_\times = h_0 \cos \iota. \quad (30)$$

where the overall GW amplitude  $h_0$  is given by

$$h_0 = \frac{4\pi^2 G}{c^4} \frac{\epsilon I_{zz} f^2}{d}, \quad (31)$$

in terms of the triaxial ellipticity  $\epsilon \equiv |I_{xx} - I_{yy}|/I_{zz}$ , and the distance  $d$ . Note that this partially fixes the gauge, namely

$$A_+ \geq |A_\times| \geq 0, \quad (32)$$

which excludes gauge-transformations (ii) and (iii) in (29). In order to fix a unique gauge also for  $\psi, \phi_0$ , we restrict the quadrant of  $\psi$  to be  $\psi \in [-\pi/4, \pi/4)$  (in accord with the TDS convention), which can always be achieved by the gauge-transformation (i), while  $\phi_0$  remains unconstrained in  $\phi_0 \in [0, 2\pi)$ .

## 2 Noise and detection statistic

### 2.1 Theoretical framework

We follow the notation of [5, 1] by denoting vectors of detector-specific quantities in boldface, i.e.  $\{\mathbf{x}\}^X = x^X$ . We can now write the explicit dependencies of the signal-model (25) on the signal-parameters as

$$\mathbf{h}(t; \mathcal{A}, \lambda) = \mathcal{A}^\mu \mathbf{h}_\mu(t; \lambda), \quad (33)$$

where we implicitly sum over amplitude-indices  $\mu, \nu \in \{1, 2, 3, 4\}$ . If the data  $x^X(t)$  measured at different detectors  $X$  consists of stationary Gaussian noise  $n^X(t)$  and a signal with parameters  $\mathcal{A}, \lambda$ , we can write

$$\mathbf{x}(t) = \mathbf{n}(t) + \mathbf{h}(t; \mathcal{A}, \lambda), \quad (34)$$

in terms of the signal-model (33). It is sometimes useful to consider the discrete-time formulation, as it more closely describes the actual measured

data, which is sampled as discrete time-steps  $t_j \equiv j \Delta t$ , namely  $x_j^X \equiv x^X(t_j)$ . The noise samples  $\{n_j^X\}$  are assumed to be drawn from a Gaussian distribution with zero mean,  $E[n_j^X] = 0$ , and covariance matrix

$$\kappa_{jl}^{XY} \equiv E[n_j^X n_l^Y] , \quad (35)$$

which allows us to write the noise probability distribution as

$$P(\mathbf{n}|\boldsymbol{\kappa}) = k e^{-\frac{1}{2}(\mathbf{n}|\mathbf{n})} , \quad (36)$$

where  $k$  is a normalization factor independent of the noise  $\mathbf{n}$ , and where we defined the discrete-time version of the multi-detector scalar product (40) as

$$(\mathbf{x}|\mathbf{y}) \equiv x_j^X \kappa_{XY}^{jl} y_l^Y , \quad (37)$$

with automatic summation over time-indices  $j, l$  and detector-indices  $X, Y$ , and  $\kappa_{XY}^{jl}$  defined as the inverse of the covariance matrix, namely

$$\kappa_{jm}^{XY} \kappa_{YZ}^{ml} = \delta_{Zj}^{Xl} . \quad (38)$$

For known functions of time  $g_j^X, h_j^X$ , and Gaussian noise  $n_j^X$  following (36), it is now easy to prove the general identity

$$\begin{aligned} E[(\mathbf{n}|\mathbf{g})(\mathbf{n}|\mathbf{h})] &= E\left[n_j^X \kappa_{XY}^{jl} g_l^Y n_m^Z \kappa_{ZV}^{mp} h_p^V\right] \\ &= g_l^Y h_p^V \kappa_{XY}^{jl} \kappa_{ZV}^{mp} \kappa_{jm}^{XZ} \\ &= g_l^Y \kappa_{YV}^{lp} h_p^V \\ &= (\mathbf{g}|\mathbf{h}) . \end{aligned} \quad (39)$$

As shown in [3] (for the single-detector case), the natural discrete-time scalar product (37), which came directly from the Gaussian probability distribution (36), leads to the well-known continuous-time formulation in the appropriate limit, namely

$$(\mathbf{x}|\mathbf{y}) \rightarrow 4 \Re \int_0^\infty \tilde{x}^X(f) S_{XY}^{-1}(f) \tilde{y}^{Y*}(f) df , \quad (40)$$

where  $\Re$  denotes the real part, and  $\tilde{x}(f)$  denotes the Fourier transformed

$$\tilde{x}(f) \equiv \int x(t) e^{-i2\pi ft} dt \approx \Delta t \sum_j x_j e^{-i2\pi f t_j} . \quad (41)$$

The matrix  $S_{XY}^{-1}$  satisfies  $S_{XY}^{-1} S^{YZ} = \delta_X^Z$ , where the (single-sided!) noise PSD matrix  $S^{XY}$  is defined as

$$S^{XY}(f) = 2 \int_{-\infty}^\infty \kappa^{XY}(\tau) e^{-i2\pi f\tau} d\tau , \quad (42)$$

in terms of the correlation matrix (assuming stationary noise)  $\kappa^{XY}(\tau) \equiv E[n^X(t+\tau)n^Y(t)]$ . In the case of uncorrelated noises between detectors, i.e.  $S^{XY} = S^X \delta^{XY}$ , the scalar product (40) reduces to a sum over single-detector scalar products, namely

$$(\mathbf{x}|\mathbf{y}) = \sum_X^{N_{\text{Det}}} (x^X|y^X) = \sum_X 4 \Re \int_0^\infty \frac{\tilde{x}^X(f) \tilde{y}^{X*}(f)}{S^X(f)} df, \quad (43)$$

where  $N_{\text{Det}}$  is the number of detectors used. Assuming  $\mathbf{x}(t)$  or  $\mathbf{y}(t)$  is a narrow-band continuous-wave signal (25) at frequency  $f_s$ , we can approximate this scalar product as

$$(\mathbf{x}|\mathbf{y}) \approx 2 \sum_X^{N_{\text{Det}}} S_X^{-1}(f_s) \int_0^T x^X(t) y^X(t) dt. \quad (44)$$

We can use the noise probability distribution (36) together with (34) to express the likelihood of observing data  $\mathbf{x} = \mathbf{n} + \mathbf{h}$  in the presence of a signal  $\mathbf{h}(t; \mathcal{A}, \lambda)$ , namely

$$P(\mathbf{x}|\mathcal{A}, \lambda, \mathbf{S}) = k e^{-\frac{1}{2}(\mathbf{x}|\mathbf{x})} e^{(\mathbf{x}|\mathbf{h}) - \frac{1}{2}(\mathbf{h}|\mathbf{h})}, \quad (45)$$

while in the noise-only case  $h_0 = 0$ , i.e.  $\mathcal{A}^\mu = 0$ , the likelihood is simply

$$P(\mathbf{x}|0, \mathbf{S}) = k e^{-\frac{1}{2}(\mathbf{x}|\mathbf{x})}. \quad (46)$$

Therefore the likelihood ratio  $\mathcal{L}(\mathbf{x}; \mathcal{A}, \lambda) \equiv P(\mathbf{x}|\mathcal{A}, \lambda) / P(\mathbf{x}|0)$  is found as

$$\begin{aligned} \log \mathcal{L}(\mathbf{x}; \mathcal{A}, \lambda) &= (\mathbf{x}|\mathbf{h}) - \frac{1}{2}(\mathbf{h}|\mathbf{h}) \\ &= \mathcal{A}^\mu x_\mu - \frac{1}{2} \mathcal{A}^\mu \mathcal{M}_{\mu\nu} \mathcal{A}^\nu, \end{aligned} \quad (47)$$

where we substituted the ‘‘JKS’’ signal factorization (33), and where we defined

$$x_\mu(\lambda) \equiv (\mathbf{x}|\mathbf{h}_\mu), \quad (48)$$

$$\mathcal{M}_{\mu\nu}(\lambda) \equiv (\mathbf{h}_\mu|\mathbf{h}_\nu) = (\partial_\mu \mathbf{h}|\partial_\nu \mathbf{h}), \quad (49)$$

defining  $\partial_\mu \equiv \partial/\partial \mathcal{A}^\mu$ . From the last expression we see that  $\mathcal{M}_{\mu\nu}$  is the Fisher matrix for the parameters  $\mathcal{A}^\mu$ . It is straightforward to analytically maximize the likelihood-ratio (47) with respect to the four amplitudes  $\mathcal{A}^\mu$ , and we obtain the so-called ‘‘ $\mathcal{F}$ -statistic’’, namely

$$\mathcal{F}(\mathbf{x}; \lambda) \equiv \max_{\mathcal{A}} \log \mathcal{L}(\mathbf{x}; \mathcal{A}, \lambda) = \frac{1}{2} x_\mu \mathcal{M}^{\mu\nu} x_\nu, \quad (50)$$



where  $\mathcal{M}^{\mu\nu} \equiv \{\mathcal{M}^{-1}\}^{\mu\nu}$ , i.e.  $\mathcal{M}_{\mu\sigma}\mathcal{M}^{\sigma\nu} = \delta_{\mu}^{\nu}$ . The maximum-likelihood (ML) estimators for the four unknown amplitudes  $\mathcal{A}^{\mu}$  are given by

$$\mathcal{A}_{\text{ML}}^{\mu} = \mathcal{M}^{\mu\nu} x_{\nu}, \quad (51)$$

and alternatively we can also express the  $\mathcal{F}$ -statistic (50) in the form

$$2\mathcal{F}(\mathbf{x}; \lambda) = \mathcal{A}_{\text{ML}}^{\mu} \mathcal{M}_{\mu\nu} \mathcal{A}_{\text{ML}}^{\nu}, \quad (52)$$

which can be interpreted as the “norm” of the ML amplitude  $\mathcal{A}_{\text{ML}}$  with respect to the “metric”  $\mathcal{M}_{\mu\nu}$  [8, 12]

## 2.2 Non-stationary, non-complete data

In practice we will be computing the power-spectra  $S_X(f)$  over shorter time-periods  $T_{\text{SFT}}$ , corresponding to the “Short Fourier Transforms” (SFT) that are used as input data to (most) CW codes. We therefore only need to assume approximately stationary noise  $S_{X\alpha}(f)$  over each SFT  $\alpha$  from detector  $X$ , allowing the noise-floor to vary from one SFT to the next. Furthermore, data might be available only for some of time during the time-span  $T$ , depending on the detector  $X$ , and we therefore base all our expressions on these SFTs as the elementary per-detector “data atoms”, writing (44) as

$$(\mathbf{x}|\mathbf{y}) \approx 2 \sum_{X=1}^{N_{\text{Det}}} \sum_{\alpha=1}^{N_{\text{SFT}}^X} S_{X\alpha}^{-1}(f) \int_0^{T_{\text{SFT}}} x_{X\alpha}(t) y_{X\alpha}(t) dt, \quad (53)$$

using the convention  $x_{X\alpha}(t) \equiv x^X(t_{X\alpha} + t)$ , where  $t_{X\alpha}$  is the start-time of the SFT  $X\alpha$ . The number of SFTs from detector  $X$  is  $N_{\text{SFT}}^X$ , i.e.

$$N_{\text{SFT}} = \sum_{X=1}^{N_{\text{Det}}} N_{\text{SFT}}^X = \sum_{X\alpha} 1, \quad (54)$$

is the total number of SFTs from all detectors. Here and in the following we use the shorthand notation

$$\sum_{X\alpha} \dots \equiv \sum_{X=1}^{N_{\text{Det}}} \sum_{\alpha=1}^{N_{\text{SFT}}^X} \dots, \quad (55)$$

to denote the sum over all used SFTs from all detectors. It will be useful to re-normalize the noise factors  $S_{X\alpha}^{-1}$  in (53), by introducing *noise weights*

$$w_{X\alpha}(f) \equiv \frac{S_{X\alpha}^{-1}(f)}{\mathcal{S}^{-1}}. \quad (56)$$

This will serve two purposes: (i) to make the weights numerically  $\sim \mathcal{O}(1)$ , and (ii) in order to allow factoring out the overall *scaling* of the scalar product with noise-floors and length of data, with the remaining factors being simple averages. Using these definitions, we can re-write the scalar product (53) as

$$\langle \mathbf{x} | \mathbf{y} \rangle \approx 2\mathcal{S}^{-1} \sum_{X\alpha} w_{X\alpha} \int_0^{T_{\text{SFT}}} x_{X\alpha}(t) y_{X\alpha}(t) dt, \quad (57)$$

which is a noise-weighted sum over single-SFT integrals. The noise-weights (56) depend on the frequency  $f$  at which they are computed, and in practice we assume  $S_{X\alpha}(f)$  to be roughly constant over a small frequency band  $\Delta f$  around the template frequency  $f_0$ . The current code (in `LALComputeMultiNoiseWeights`) defines the weights in terms of the *arithmetic mean* of the PSD over  $\Delta f$  of the input SFTs, i.e.

$$w_{X\alpha}(f_0) \approx \frac{\langle S_{X\alpha}(f) \rangle_{f_0 \pm \Delta f/2}^{-1}}{\mathcal{S}^{-1}}. \quad (58)$$

The normalization constant  $\mathcal{S}^{-1}$  is in principle arbitrary and drops out from any physically meaningful result. For practical purposes, however, we choose it in such a way to achieve (i) and (ii) mentioned above, namely

$$\sum_{X\alpha} w_{X\alpha} = N_{\text{SFT}}, \quad \text{therefore} \quad (59)$$

$$\mathcal{S}^{-1} \equiv \frac{1}{N_{\text{SFT}}} \sum_{X\alpha} S_{X\alpha}^{-1}. \quad (60)$$

Using this convention,  $\mathcal{S}$  is defined as the *harmonic mean* over the per-SFT noise PSDs  $S_{X\alpha}$  over all SFTs  $\alpha$  from all detectors  $X$ . These weights have the property that  $N_{\text{SFT}}^{-1} \sum_{X\alpha} w_{X\alpha} = 1$ , and so we can conveniently define a *total noise-weighted average*  $\langle x y \rangle_w$ , namely<sup>2</sup>

$$\langle x y \rangle_w \equiv \frac{1}{N_{\text{SFT}}} \sum_{X\alpha} w_{X\alpha} \langle x_{X\alpha} y_{X\alpha} \rangle_t, \quad (61)$$

in terms of single-SFT time-averages  $\langle Z_{X\alpha} \rangle_t$  of a function  $Z_{X\alpha}(t)$  of time and detector, defined as

$$\langle Z_{X\alpha} \rangle_t \equiv \frac{1}{T_{\text{SFT}}} \int_0^{T_{\text{SFT}}} Z_{X\alpha}(t) dt. \quad (62)$$

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<sup>2</sup>Note that our definition of  $\mathcal{S}^{-1}$  and averaging operator  $\langle \dots \rangle_w$  here differ from the conventions used in [6], which are less symmetric in time and detectors, and less suitable for generalization to varying noise-floors.

Using this, the scalar product (57) can now be expressed as

$$(\mathbf{x}|\mathbf{y}) \approx 2\mathcal{S}^{-1}T_{\text{data}} \langle xy \rangle_w, \quad (63)$$

where

$$T_{\text{data}} \equiv N_{\text{SFT}} T_{\text{SFT}} \quad (64)$$

is the total time length of data used.

The scalar products involved in the  $\mathcal{F}$ -statistic contain slowly-varying (diurnal) antenna-pattern functions  $\{a(t), b(t)\}$ , and phase-functions  $\{\sin \phi(t), \cos \phi(t)\}$  that are oscillatory on short timescales  $1/f \ll T_{\text{SFT}}$ . Using these properties, the  $4 \times 4$  matrix  $\mathcal{M}_{\mu\nu}$  defined in Eq. (49), namely

$$\mathcal{M}_{\mu\nu} \equiv (\mathbf{h}_\mu|\mathbf{h}_\nu) = \mathcal{S}^{-1}T_{\text{data}} m_{\mu\nu}, \quad (65)$$

can be approximated to yield the block-form

$$m_{\mu\nu} = 2\langle h_\mu h_\nu \rangle_w \approx \begin{pmatrix} A & C & 0 & E \\ C & B & -E & 0 \\ 0 & -E & A & C \\ E & 0 & C & B \end{pmatrix}, \quad (66)$$

with the 3 (+1 in the RAA case) independent components

$$A \equiv \langle |a|^2 \rangle_w, \quad B \equiv \langle |b|^2 \rangle_w, \quad C \equiv \Re \langle a^* b \rangle_w, \quad E \equiv \Im \langle a^* b \rangle_w, \quad (67)$$

where  $E = 0$  in the LWL limit. We further define the determinant-factor  $D \equiv AB - C^2 - E^2$ , such that  $\det m = D^2$ .

Introducing the complex matched filters

$$\begin{aligned} x_a &\equiv x_1 - ix_3 = (\mathbf{x}|\mathbf{h}_a), \\ x_b &\equiv x_2 - ix_4 = (\mathbf{x}|\mathbf{h}_b), \end{aligned} \quad (68)$$

in terms of the complex basis (28), and using (65), we can now write the  $\mathcal{F}$ -statistic (50) more explicitly as

$$2\mathcal{F} = \frac{D^{-1}}{\mathcal{S}^{-1}T_{\text{data}}} [B|x_a|^2 + A|x_b|^2 - 2C\Re(x_a^* x_b) - 2E\Im(x_a^* x_b)]. \quad (69)$$

### 2.3 $\mathcal{F}$ -statistic of perfectly matched signal

Let us assume there is a signal  $\mathbf{s}(t)$  in the data that is perfectly matched by the search-template, i.e.

$$\begin{aligned} \mathbf{x}(t) &= \mathbf{n}(t) + \mathbf{s}(t), \quad \text{where} \\ \mathbf{s}(t) &= \mathbf{h}(t; \mathcal{A}_s, \lambda_s) = \mathcal{A}_s^\mu \mathbf{h}_\mu(t; \lambda_s), \end{aligned} \quad (70)$$

and so the four amplitude-components  $x_\mu$ , defined in (48), are

$$x_\mu(\mathcal{A}_s, \lambda_s) = n_\mu(\lambda_s) + s_\mu(\mathcal{A}_s, \lambda_s), \quad (71)$$

where  $n_\mu \equiv (\mathbf{n}|\mathbf{h}_\mu)$  and

$$s_\mu \equiv (\mathbf{s}|\mathbf{h}_\mu) = \mathcal{A}_s^\nu \mathcal{M}_{\nu\mu}(\lambda_s). \quad (72)$$

One can show the following identities for zero-mean Gaussian noise  $\mathbf{n}$ :

$$E[n_\mu] = 0, \quad \text{and} \quad E[n_\mu n_\nu] = \mathcal{M}_{\mu\nu}, \quad (73)$$

where in the second equation we used (39). This results in

$$E[x_\mu] = s_\mu, \quad \text{and} \quad E[x_\mu x_\nu] = \mathcal{M}_{\mu\nu} + s_\mu s_\nu, \quad (74)$$

which shows that the four random variables  $x_\mu$  have means  $s_\mu$  and covariance  $\mathcal{M}_{\mu\nu}$  (independent of the signal strength). By applying these relations to Eq. (50), we find the expectation of  $2\mathcal{F}$  in the perfectly-matched case as

$$E[2\mathcal{F}] = 4 + \rho^2(0), \quad (75)$$

where we defined the “optimal” signal-to-noise ratio (SNR)  $\rho(0)$  as

$$\rho^2(0) \equiv s_\mu \mathcal{M}^{\mu\nu} s_\nu = \mathcal{A}_s^\mu \mathcal{M}_{\mu\nu} \mathcal{A}_s^\nu = (\mathbf{s}|\mathbf{s}). \quad (76)$$

Combining (26) and (65), (66) this can be written<sup>3</sup> more explicitly as

$$\rho^2(0) = h_0^2 (\alpha_1 A + \alpha_2 B + 2\alpha_3 C + 2\alpha_4 E) \mathcal{S}^{-1} T_{\text{data}}, \quad (77)$$

where the functions  $\alpha_i(\eta, \psi)$  are defined as (with  $\eta \equiv \cos \iota$ ):

$$\alpha_1(\eta, \psi) \equiv (\hat{\mathcal{A}}^1)^2 + (\hat{\mathcal{A}}^3)^2 = \frac{1}{4}(1 + \eta^2)^2 \cos^2 2\psi + \eta^2 \sin^2 2\psi, \quad (78)$$

$$\alpha_2(\eta, \psi) \equiv (\hat{\mathcal{A}}^2)^2 + (\hat{\mathcal{A}}^4)^2 = \frac{1}{4}(1 + \eta^2)^2 \sin^2 2\psi + \eta^2 \cos^2 2\psi, \quad (79)$$

$$\alpha_3(\eta, \psi) \equiv \hat{\mathcal{A}}^1 \hat{\mathcal{A}}^2 + \hat{\mathcal{A}}^3 \hat{\mathcal{A}}^4 = \frac{1}{4}(1 - \eta^2)^2 \sin 2\psi \cos 2\psi, \quad (80)$$

$$\alpha_4(\eta, \psi) \equiv \hat{\mathcal{A}}^1 \hat{\mathcal{A}}^4 - \hat{\mathcal{A}}^2 \hat{\mathcal{A}}^3 = \frac{1}{2} \eta (1 + \eta^2). \quad (81)$$

using the re-scaled amplitude parameters  $\hat{\mathcal{A}}^\mu \equiv \mathcal{A}^\mu/h_0$ .

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<sup>3</sup>The difference to Eq. (68) of [6] is the use of single-sided noise PSD, and the different definitions of  $\mathcal{S}^{-1}$  and averaging operator

## 2.4 Average SNR<sup>2</sup>

It is often useful to compute averaged quantities over the amplitude parameters  $\{\cos \iota, \psi\}$  and sky-position  $\vec{n}$ . Averaging a quantity  $Z$  over  $\{\cos \iota, \psi\}$  with isotropic priors on the source-orientation, which translates into uniform priors [7] over  $\cos \iota$  and  $\psi$ , namely

$$\langle Z \rangle_{\cos \iota, \psi} \equiv \frac{1}{2} \int_{-1}^1 d \cos \iota \frac{1}{\pi/2} \int_{-\pi/4}^{\pi/4} d\psi Z(\cos \iota, \psi), \quad (82)$$

yields

$$\begin{aligned} \langle \alpha_1 \rangle_{\cos \iota, \psi} &= \langle \alpha_2 \rangle_{\cos \iota, \psi} = \frac{2}{5}, \\ \langle \alpha_3 \rangle_{\cos \iota, \psi} &= \langle \alpha_4 \rangle_{\cos \iota, \psi} = 0. \end{aligned} \quad (83)$$

It is also useful to consider the averages of  $\alpha_i$  over  $\psi$  and  $\cos \iota$  separately, as first noted in [10], for which we obtain

$$\begin{aligned} \langle \alpha_1 \rangle_{\psi} &= \langle \alpha_2 \rangle_{\psi} = \frac{1}{8} (\eta^4 + 6\eta^2 + 1), \\ \langle \alpha_3 \rangle_{\psi} &= 0, \quad \langle \alpha_4 \rangle_{\psi} = \alpha_4. \end{aligned} \quad (84)$$

and

$$\begin{aligned} \langle \alpha_1 \rangle_{\cos \iota} &= \frac{1}{15} (6 + \cos 4\psi), \quad \langle \alpha_2 \rangle_{\cos \iota} = \frac{1}{15} (6 - \cos 4\psi), \\ \langle \alpha_3 \rangle_{\cos \iota} &= \frac{1}{15} \sin 4\psi, \quad \langle \alpha_4 \rangle_{\cos \iota} = 0. \end{aligned} \quad (85)$$

The sky-averages of  $A, B, C, E$  are a little more involved. From (67) we see that these antenna-pattern coefficients are time-averages and noise-weighted detector averages of  $|a|^2$ ,  $|b|^2$ , and  $a^* b$  respectively, with the antenna-pattern functions  $a(t; \vec{n})$  and  $b(t; \vec{n})$  defined in (15). We can therefore change the order of isotropic sky-averaging, defined as

$$\langle Z \rangle_{\vec{n}} \equiv \frac{1}{4\pi} \int_0^{2\pi} d\alpha \int_{-1}^1 Z(\vec{n}) d \sin \delta, \quad (86)$$

with (per-detector) time-averaging in (61), and we therefore obtain

$$\langle A^X \rangle_{\vec{n}} = \langle \langle |a_X|^2 \rangle_t \rangle_{\vec{n}} = \langle \langle |a_X|^2 \rangle_{\vec{n}} \rangle_t = \langle |a_X|^2 \rangle_{\vec{n}}, \quad (87)$$

(and similarly for  $B, C, E$ ) where in the last step we used the fact that the sky-averaged antenna-pattern function does not depend on time. This can

be seen by noting that the time-dependency is a diurnal rotation around  $\hat{z}$ , resulting in a pure translation in right-ascension coordinate, with declination and orientation with respect to polarization basis unchanged.

However, while the all-sky antenna-pattern averages are independent of time, they *do* depend on the orientation of the detector arms  $\{\hat{l}_1, \hat{l}_2\}$  with respect to the polarization basis of Eq. (6), and are therefore different for different detector orientations.

**Long-wavelength limit expressions:** Note that for simplicity in the following we assume the long-wavelength limit (LWL) expression for the detector tensor Eq. (10), which implies that  $a(t), b(t)$  are real-valued and  $E = 0$ . It is convenient to re-parametrize the detector tensor in terms of the two orthogonal insensitive directions (c.f. [10]) of the detector, namely

$$\hat{m}_1 \equiv \frac{1}{2 \sin(\zeta/2)} (\hat{l}_1 - \hat{l}_2), \quad \hat{m}_2 \equiv \frac{1}{2 \cos(\zeta/2)} (\hat{l}_1 + \hat{l}_2), \quad (88)$$

where  $\zeta$  is the opening angle between the two arms ( $\hat{l}_1 \cdot \hat{l}_2 = \cos \zeta$ ). This yields the expression for the detector tensor

$$\underline{d} = \frac{\sin \zeta}{2} (\hat{m}_1 \otimes \hat{m}_2 + \hat{m}_2 \otimes \hat{m}_1). \quad (89)$$

Without loss of generality, for the sky-average we can express the general detector-arm orientation as

$$\hat{m}_1 = (0, \cos \psi_0, \sin \psi_0), \quad \hat{m}_2 = (0, \sin \psi_0, -\cos \psi_0), \quad (90)$$

which is one (equivalent) instance of the class of configurations  $(\zeta, \psi_0)$ , where the angle  $\psi_0$  is the rotation counter-clockwise around the detector zenith-direction  $\hat{n}_0 = \hat{l}_1 \times \hat{l}_2 = \sin \zeta (\hat{m}_1 \times \hat{m}_2)$ , from the polarization basis vector  $\hat{\xi}_0(\hat{n}_0)$  to  $\hat{m}_1$ , for  $\hat{\xi}_0 = \hat{n}_0 \times \hat{z} / |\hat{n}_0 \times \hat{z}|$ , with  $\hat{z}$  pointing north. Inserting these expressions into the definitions (15) for  $a^X, b^X$ , we can obtain the explicit sky-averaged expressions

$$\begin{aligned} \langle A^X \rangle_{\bar{n}} &= \langle a_X^2 \rangle_{\bar{n}} = \frac{2}{5} \sin^2 \zeta - \langle B^X \rangle_{\bar{n}}, \\ \langle B^X \rangle_{\bar{n}} &= \langle b_X^2 \rangle_{\bar{n}} = \sin^2 \zeta \frac{7 \cos(4\psi_0) + 9}{48}, \\ \langle C^X \rangle_{\bar{n}} &= \langle a_X b_X \rangle_{\bar{n}} = 0. \end{aligned} \quad (91)$$

such that  $\langle A^X \rangle_{\bar{n}} + \langle B^X \rangle_{\bar{n}} = \frac{2}{5} \sin^2 \zeta$ .

Equipped with these averages, we can now obtain from (77)

$$\langle \rho^2 \rangle_{\cos \iota, \psi} = \frac{2}{5} h_0^2 (A + B) \mathcal{S}^{-1} T_{\text{data}}, \quad (92)$$

$$\langle \rho^2 \rangle_{\vec{n}} = h_0^2 (\alpha_1 \langle A^X \rangle_{\vec{n}, w} + \alpha_2 \langle B^X \rangle_{\vec{n}, w}) \mathcal{S}^{-1} T_{\text{data}}, \quad (93)$$

$$\langle \rho^2 \rangle_{\vec{n}, \cos \iota, \psi} = \frac{4}{25} h_0^2 \mathcal{S}^{-1} T_{\text{data}} \sin^2 \zeta, \quad (94)$$

in agreement with Eq.(93) in [4]. Furthermore, a selective average over sky and  $\psi$ , as first considered in [10], yields

$$\langle \rho^2 \rangle_{\vec{n}, \psi} = \frac{1}{20} h_0^2 (\eta^4 + 6\eta^2 + 1) \mathcal{S}^{-1} T_{\text{data}} \sin^2 \zeta. \quad (95)$$

It is sometimes convenient to express the instantaneous “strength” of a signal in the detectors, independently of the observation time and noise floor, and following [2] we define  $h_{\text{rms}}$  as the root-mean-square (rms) of the signal strain, averaged over time and detectors, i.e.

$$h_{\text{rms}}^2 \equiv \langle s^2 \rangle_w = \mathcal{A}_s^\mu \langle h_\mu h_\nu \rangle_w \mathcal{A}_s^\nu = \frac{1}{2} \mathcal{A}_s^\mu m_{\mu\nu} \mathcal{A}_s^\nu, \quad (96)$$

in terms of the antenna-pattern matrix  $m_{\mu\nu}$  defined in (66). Using this definition and (65), the optimal SNR (76) can now also be written as

$$\rho^2 = 2\mathcal{S}^{-1} T_{\text{data}} h_{\text{rms}}^2, \quad (97)$$

and comparing this to (77) we obtain

$$h_{\text{rms}}^2 = \frac{1}{2} h_0^2 (\alpha_1 A + \alpha_2 B + 2\alpha_3 C). \quad (98)$$

Averaging this over all sky-positions  $\vec{n}$  and polarization angles  $\cos \iota, \psi$  at fixed amplitude  $h_0$ , we find

$$\langle h_{\text{rms}}^2 \rangle_{\cos \iota, \psi, \vec{n}} = \frac{2}{25} h_0^2 \sin^2 \zeta. \quad (99)$$

### 3 Parameter estimation of the signal

#### 3.1 Estimating amplitude parameters $\{h_0, \cos \iota, \psi, \phi_0\}$

From the expression (51) for the maximum-likelihood amplitudes  $\mathcal{A}_{\text{ML}}^\mu$  in terms of the measured  $x_\mu$ , we can infer the signal-parameters  $A_+, A_\times$  (or equivalently  $h_0, \cos \iota$ ) and  $\psi, \phi_0$ , by using (26) and (30), mostly following

Yousuke's notes. We want to invert the four relations  $\mathcal{A}^\mu(h_0, \cos \iota, \psi, \phi_0)$  in Eq. (26), and we start by computing the two quantities

$$A_s^2 \equiv \sum_{\mu=1}^4 (\mathcal{A}^\mu)^2 = A_+^2 + A_\times^2, \quad (100)$$

$$D_a \equiv \mathcal{A}^1 \mathcal{A}^4 - \mathcal{A}^2 \mathcal{A}^3 = A_+ A_\times, \quad (101)$$

which can easily be solved for  $A_+$ ,  $A_\times$ , namely

$$2A_{+,\times}^2 = A_s^2 \pm \sqrt{A_s^4 - 4D_a^2}, \quad (102)$$

where our convention here is  $|A_+| \geq |A_\times|$ , cf. (30), and therefore the '+' solution is  $A_+$ , and the '-' is  $A_\times$ . The sign of  $A_+$  is always positive by convention (30), while the sign of  $A_\times$  is given by the sign of  $D_a$ , as can be seen from (101). Note that the discriminant in (102) is also expressible as

$$\text{disc} \equiv \sqrt{A_s^4 - 4D_a^2} = A_+^2 - A_\times^2 \geq 0. \quad (103)$$

Having computed  $A_+$ ,  $A_\times$ , we can now also obtain  $\psi$  and  $\phi_0$ , namely defining  $\beta \equiv A_\times/A_+$ , and

$$b_1 \equiv \mathcal{A}^4 - \beta \mathcal{A}^1, \quad (104)$$

$$b_2 \equiv \mathcal{A}^3 + \beta \mathcal{A}^2, \quad (105)$$

$$b_3 \equiv \beta \mathcal{A}^4 - \mathcal{A}^1, \quad (106)$$

we easily find

$$\psi = \frac{1}{2} \text{atan} \left( \frac{b_1}{b_2} \right). \quad (107)$$

and

$$\phi_0 = \text{atan} \left( \frac{b_2}{b_3} \right). \quad (108)$$

Note that there is still an overall sign-ambiguity in the amplitudes  $\mathcal{A}^\mu$ , which can be determined as follows: compute a 'reconstructed'  $\mathcal{A}_r^1$  from (26) using the estimates  $A_{+,\times}$  and  $\psi, \phi_0$ , and compare its sign to the original estimate  $\mathcal{A}^1$  of (135). If the sign differs, the correct solution is simply found by replacing  $\phi_0 \rightarrow \phi_0 + \pi$ .

Converting  $A_+, A_\times$  into  $h_0$  and  $\mu \equiv \cos \iota$  is done by solving (30), which yields

$$h_0 = A_+ + \sqrt{A_+^2 - A_\times^2}, \quad (109)$$

where we only kept the '+' solution, as we must have  $h_0 \geq A_+$  (which can be seen from (30)). Finally,  $\mu = \cos \iota$  is simply given by  $\cos \iota = A_\times/h_0$ .



### 3.2 Errors in amplitude-parameter estimation

Let us define the error  $\Delta\mathcal{A}^\mu$  in maximum-likelihood parameter estimation on the four amplitude  $\mathcal{A}^\mu$  simply as

$$\Delta\mathcal{A}^\mu \equiv \mathcal{A}_{\text{ML}}^\mu - \mathcal{A}_s^\mu. \quad (110)$$

Given (51), (71) and (72), we have

$$\mathcal{A}_{\text{ML}}^\mu = \mathcal{M}^{\mu\nu} n_\nu + \mathcal{A}_s^\nu, \quad (111)$$

and therefore

$$\Delta\mathcal{A}^\mu = \mathcal{M}^{\mu\nu} n_\nu, \quad (112)$$

and so we directly obtain using (73)

$$E[\mathcal{A}_{\text{ML}}^\mu] = \mathcal{A}_s^\mu, \quad \text{i.e.} \quad E[\Delta\mathcal{A}^\mu] = 0, \quad (113)$$

namely the ML estimators for the  $\mathcal{A}^\mu$  are *unbiased*. Furthermore, the covariance matrix of the errors  $\Delta\mathcal{A}^\mu$  is found as

$$E[\Delta\mathcal{A}^\mu \Delta\mathcal{A}^\nu] = \mathcal{M}^{\mu\nu}, \quad (114)$$

which corresponds to the Cramér-Rao bound, where  $\mathcal{M}^{\mu\nu}$  is the inverse of the Fisher matrix (49). Note that we have not made any assumptions about the errors  $\Delta\mathcal{A}^\mu$  being “small”, the Fisher-matrix relation (114) is strictly true here for any deviations and SNR, *provided* the  $\mathcal{A}_{\text{ML}}^\mu$  were measured at exactly the right signal Doppler location  $\lambda_s$ , such that  $\mathcal{M}_{\mu\nu} = \mathcal{M}_{\mu\nu}(\lambda_s)$ . Any parameter-estimation error in  $\lambda$  would complicate the picture, which is why these error-estimates strictly only apply in a perfectly-matched (“targeted”) search case.

Let us now consider arbitrary functions  $f_i(\mathcal{A}^\mu)$  of the four amplitudes  $\mathcal{A}^\mu$ , where for *small* errors  $df_i$  we have

$$df_i = \partial_\mu f_i d\mathcal{A}^\mu, \quad (115)$$

and therefore we can find the error covariances

$$E[df_i df_j] = \partial_\mu f_i \mathcal{M}^{\mu\nu} \partial_\nu f_j. \quad (116)$$

We can consider different more “physical” amplitude-parameter coordinates such as  $\mathcal{A}^{\hat{i}} \equiv (A_+, A_\times, \phi_0, \psi)$  or  $\mathcal{A}^i \equiv (h_0, \cos \iota, \phi_0, \psi)$ . From (26) one easily obtains the explicit Jacobian

$$J^{\mu}_{\hat{i}} \equiv \frac{\partial \mathcal{A}^\mu}{\partial \mathcal{A}^{\hat{i}}} = \begin{pmatrix} \cos \phi_0 \cos 2\psi & -\sin \phi_0 \sin 2\psi & \mathcal{A}^3 & -2 \mathcal{A}^2 \\ \cos \phi_0 \sin 2\psi & \sin \phi_0 \cos 2\psi & \mathcal{A}^4 & 2 \mathcal{A}^1 \\ -\sin \phi_0 \cos 2\psi & -\cos \phi_0 \sin 2\psi & -\mathcal{A}^1 & -2 \mathcal{A}^4 \\ -\sin \phi_0 \sin 2\psi & \cos \phi_0 \cos 2\psi & -\mathcal{A}^2 & 2 \mathcal{A}^3 \end{pmatrix} \quad (117)$$

and by (numerical) inversion we can obtain  $\partial_\mu \mathcal{A}^{\hat{i}} = J^{-1\hat{i}}{}_\mu$ . We therefore can compute the covariance matrix of errors  $d\mathcal{A}^{\hat{i}}$  from (116), namely

$$E[d\mathcal{A}^{\hat{i}} d\mathcal{A}^{\hat{j}}] = J^{-1\hat{i}}{}_\mu J^{-1\hat{j}}{}_\nu \mathcal{M}^{\mu\nu}. \quad (118)$$

Similarly, for the choice of output-variables  $\mathcal{A}^i$ , using (30) we find

$$\frac{\partial \mathcal{A}^\mu}{\partial h_0} = \frac{\mathcal{A}^\mu}{h_0}, \quad \frac{\partial \mathcal{A}^\mu}{\partial \cos \iota} = B^\mu, \quad (119)$$

where we defined

$$B^\mu \equiv A_\times \frac{\partial \mathcal{A}^\mu}{\partial A_\times} + h_0 \frac{\partial \mathcal{A}^\mu}{\partial A_+} = \{\mathcal{A}^\mu | \text{replace } A_\times \mapsto h_0, A_+ \mapsto A_\times\}, \quad (120)$$

and so we obtain the corresponding Jacobian

$$J^\mu{}_i \equiv \frac{\partial \mathcal{A}^\mu}{\partial \mathcal{A}^i} = \begin{pmatrix} \mathcal{A}^1/h_0 & B^1 & \mathcal{A}^3 & -2\mathcal{A}^2 \\ \mathcal{A}^2/h_0 & B^2 & \mathcal{A}^4 & 2\mathcal{A}^1 \\ \mathcal{A}^3/h_0 & B^3 & -\mathcal{A}^1 & -2\mathcal{A}^4 \\ \mathcal{A}^4/h_0 & B^4 & -\mathcal{A}^2 & 2\mathcal{A}^3 \end{pmatrix} \quad (121)$$

and we can obtain the covariance matrix of small errors  $d\mathcal{A}^i$  as

$$E[d\mathcal{A}^i d\mathcal{A}^j] = J^{-1i}{}_\mu J^{-1j}{}_\nu \mathcal{M}^{\mu\nu}. \quad (122)$$

Note, however, that both (118) and (122) are only valid in the limit of *small* errors  $d$  (i.e. the high-SNR limit), and are potentially subject to singularities in the coordinate transformations, i.e. (117) (121). The formulation (114) in ‘‘canonical’’ coordinates  $\mathcal{A}^\mu$  is generally true at any SNR and is always well-defined.

## 4 Practical computation in CFS\_v2

### 4.1 Data normalization and antenna weighting

The expectation value of the  $\mathcal{F}$ -statistic is  $E[2\mathcal{F}] = 4 + \text{SNR}^2$ . For practical and numerical convenience, we want to make all quantities involved in computing  $\mathcal{F}$  of order  $\mathcal{O}(1)$ . This is already the case for the antenna-pattern functions  $\{A, B, C\}$ , defined in (67). However, the scale of the input data  $x^X(t)$  is vastly different, namely from the Wiener-Khinchin theorem we can estimate<sup>4</sup> the (single-sided) PSDs  $S_{X\alpha}(f)$  as

$$E[|\tilde{x}_{X\alpha}(f)|^2] \approx \frac{1}{2} T_{\text{SFT}} S_{X\alpha}(f) \sim \mathcal{O}(10^{-44} \text{s}^2), \quad (123)$$

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<sup>4</sup>This is the basis for estimating the noise PSD in the function `LALNormalizeSFT()`.

where  $\tilde{x}_{X\alpha}(f)$  is the ‘‘Short Fourier transform’’ (SFT), defined as

$$\tilde{x}_{X\alpha}(f) = \int_0^{T_{\text{SFT}}} x_{X\alpha}(t) e^{-i2\pi ft} dt = T_{\text{SFT}} \langle x_{X\alpha}(t) e^{-i2\pi ft} \rangle_t. \quad (124)$$

Therefore, if we re-normalize<sup>5</sup> the data as (`LALNormalizeMultiSFTVect()`)<sup>6</sup>:

$$\tilde{y}_{X\alpha}(f) \equiv \frac{\tilde{x}_{X\alpha}(f)}{\sqrt{\frac{1}{2}T_{\text{SFT}}S_{X\alpha}(f)}} \approx \frac{\tilde{x}_{X\alpha}(f)}{\sqrt{E[|\tilde{x}_{X\alpha}(f)|^2]}}, \quad (125)$$

then  $E[|\tilde{y}_{X\alpha}(f)|^2] = 1$  and therefore  $\tilde{y}_{X\alpha} \sim \mathcal{O}(1)$ . Note, however, that in practice we *estimate*  $E[|\tilde{x}_{X\alpha}(f)|^2]$  from the median of a finite number of neighboring bins. The fluctuations in this noise-floor estimator introduce a bias in (125), namely  $E[|\tilde{y}_{X\alpha}(f)|^2] \gtrsim 1$ , resulting in a bias in  $\mathcal{F}$ , namely  $E[2\mathcal{F}] \gtrsim 4$  in pure noise. Substituting (125) into (68) using the scalar product (57), we find

$$x_a = \sqrt{2\mathcal{S}^{-1}T_{\text{SFT}}} \sum_{X\alpha} \sqrt{w_{X\alpha}} \int_0^{T_{\text{SFT}}} y_{X\alpha}(t) a_{X\alpha}(t) e^{-i\phi_{X\alpha}(t)} dt, \quad (126)$$

and similarly for  $x_b$ . Furthermore, expanding (67) into

$$A \equiv \langle a^2 \rangle_w = \frac{1}{N_{\text{SFT}}} \sum_{X\alpha} w_{X\alpha} \langle a_{X\alpha}^2 \rangle_t, \quad (127)$$

we see that we can completely absorb the noise-weights  $w_{X\alpha}$  into  $\{a_{X\alpha}(t), b_{X\alpha}(t)\}$ , namely by defining *noise-weighted* antenna-pattern functions

$$\hat{a}_{X\alpha}(t) \equiv \sqrt{w_{X\alpha}} a_{X\alpha}(t), \quad \hat{b}_{X\alpha}(t) \equiv \sqrt{w_{X\alpha}} b_{X\alpha}(t). \quad (128)$$

We can now write

$$x_{\{a,b\}} = \sqrt{2\mathcal{S}^{-1}T_{\text{SFT}}} F_{\{a,b\}}, \quad (129)$$

$$\{A, B, C\} = \frac{1}{N_{\text{SFT}}} \{\hat{A}, \hat{B}, \hat{C}\}, \quad (130)$$

<sup>5</sup>In the special `--SignalOnly` case the `CFS_v2` code does not try to normalize the data and instead *assumes* the (single-sided) noise-power to be  $S_{X\alpha} = 1$ . The ‘‘missing’’ normalization-factor of  $\sqrt{T_{\text{SFT}}/2}$  is then applied to  $F_{\{a,b\}}$  a-posteriori.

<sup>6</sup>There is a small inconsistency here: in the definition of the noise-weights (58), we used the frequency-averaged  $\langle S_{X\alpha} \rangle_{\Delta f}$  over the Band  $\Delta f$  of the SFT, while in the data-normalization (125) we use the per-bin values of  $S_{X\alpha}(f)$ .

introducing the quantities  $F_{\{a,b\}}$  and  $\{\widehat{A}, \widehat{B}, \widehat{C}\}$  that are used in the `CFS_v2` code, and which are defined as

$$F_{\{a,b\}} \equiv \sum_{X\alpha} F_{\{a,b\}}^{X\alpha}, \quad (131)$$

$$F_a^{X\alpha} \equiv \int_0^{T_{\text{SFT}}} y_{X\alpha}(t) \widehat{a}_{X\alpha}(t) e^{-i\phi_{X\alpha}(t)} dt, \quad F_b^{X\alpha} = \dots (\widehat{a} \mapsto \widehat{b}) \quad (132)$$

$$\widehat{A} \equiv \sum_{X\alpha} \langle \widehat{a}_{X\alpha}^2 \rangle_t, \quad \widehat{B} \equiv \sum_{X\alpha} \langle \widehat{b}_{X\alpha}^2 \rangle_t, \quad \widehat{C} \equiv \sum_{X\alpha} \langle \widehat{a}_{X\alpha} \widehat{b}_{X\alpha} \rangle_t, \quad (133)$$

Inserting (129)(130) into (69), we obtain

$$2\mathcal{F} = \frac{2}{\widehat{D}} \left[ \widehat{B} |F_a|^2 + \widehat{A} |F_b|^2 - 2\widehat{C} \Re(F_a F_b^*) \right], \quad (134)$$

with  $\widehat{D} \equiv \widehat{A}\widehat{B} - \widehat{C}^2$ . We can express the maximum-likelihood estimators (51) for the amplitudes  $\mathcal{A}^\mu$  explicitly as

$$\mathcal{A}_{\text{ML}}^\mu = \mathcal{M}^{\mu\nu} x_\nu = \frac{\sqrt{2}\widehat{D}^{-1}}{\sqrt{\mathcal{S}^{-1}T_{\text{SFT}}}} \begin{pmatrix} \widehat{B} F_a^{\Re} - \widehat{C} F_b^{\Re} \\ -\widehat{C} F_a^{\Re} + \widehat{A} F_b^{\Re} \\ -\widehat{B} F_a^{\Im} + \widehat{C} F_b^{\Im} \\ \widehat{C} F_a^{\Im} - \widehat{A} F_b^{\Im} \end{pmatrix}, \quad (135)$$

with  $F_{\{a,b\}}^{\Re} \equiv \Re F_{\{a,b\}}$ , and  $F_{\{a,b\}}^{\Im} \equiv \Im F_{\{a,b\}}$ .

We see from (131)–(134) that the  $\mathcal{F}$ -statistic is computed completely from the set of per-SFT “ $\mathcal{F}$ -atoms”

$$\{F_{\{a,b\}}^{X\alpha}, \langle \widehat{a}_{X\alpha}^2 \rangle_t, \langle \widehat{b}_{X\alpha}^2 \rangle_t, \langle \widehat{a}_{X\alpha} \widehat{b}_{X\alpha} \rangle_t\}. \quad (136)$$

These “ $\mathcal{F}$ -atoms” are also the primary input to `CFS_v2` for the transient-CW search over different start-times and durations, as described in [9].

## 4.2 The Williams-Schutz approximation (“LALDemod”)

This section is originally based on Xavie’s LALDemod-notes<sup>7</sup>, and the method is largely based on [13].

With the convention introduced in (53), the (normalized) data time-series corresponding to an SFT  $X\alpha$  of duration  $T_{\text{SFT}}$  is written as

$$y_{X\alpha j} = y_{X\alpha}(t_j) = y(t_{X\alpha} + j\Delta t), \quad (137)$$

<sup>7</sup> [www.lsc-group.phys.uwm.edu/~siemens/demod.pdf](http://www.lsc-group.phys.uwm.edu/~siemens/demod.pdf)

where  $j = 0, \dots, N - 1$  such that  $T_{\text{SFT}} = N\Delta t$ , and  $t_{X\alpha}$  is the start-time of the SFT  $X\alpha$ . As noted above, all components of  $\mathcal{F}$  entering (134), namely  $F_{\{a,b\}}$  and  $\{\widehat{A}, \widehat{B}, \widehat{C}\}$  are *sums* over per-SFT “ $\mathcal{F}$ -atoms” (136). Here we focus on the calculation of the atoms  $F_{\{a,b\}}^{X\alpha}$ , and in order to simplify the notation we drop the SFT-index  $X\alpha$  from most of the following expressions, which refer to quantities evaluated for a single SFT  $X\alpha$ . The frequency-domain SFT data is computed from the discretized version of (124), namely

$$\tilde{y}_k \equiv \Delta t \sum_{j=0}^{N-1} y_j e^{-i2\pi kj/N}, \quad (138)$$

which is exactly what is stored in an SFT-file according to the “SFT-v2” specification (LIGO-T040164-01-Z), where in practice we only store the first  $\lfloor N/2 \rfloor$  frequency-bins, as for real  $y_j$  we have  $\tilde{y}_{N-k|N} = \tilde{y}_k^*$ . The inverse operation to (138) is

$$y_j = \Delta f \sum_{k=0}^{N-1} \tilde{y}_k e^{i2\pi kj/N}. \quad (139)$$

We write the discretized version of (132) as

$$F_a^{X\alpha} = \Delta t \sum_{j=0}^{N-1} y_j \hat{a}_j e^{-i2\pi \varphi_j}, \quad (140)$$

where we defined  $2\pi\varphi_j \equiv \phi(t_j)$  for later convenience.

The typical SFT-duration (e.g. half an hour) is chosen to be short compared to the variability of the signal, and so we can approximate the antenna-pattern functions as nearly constant over this period. Writing the SFT-midpoint as  $t_{\frac{1}{2}} \equiv T_{\text{SFT}}/2$ , we approximate  $\hat{a}_j \approx \hat{a} \equiv \hat{a}(t_{\frac{1}{2}})$ . Using this and the inverse DFT (139), we can write (140) as

$$F_a^{X\alpha} \approx \hat{a} \Delta f \Delta t \sum_{j=0}^{N-1} e^{-i2\pi \varphi_j} \sum_{k=0}^{N-1} \tilde{y}_k e^{i2\pi jk/N}. \quad (141)$$

The phase-evolution of a typical continuous pulsar-signal is dominated by the linear term  $\phi(t) \approx 2\pi f t$ , and we approximate it by a first-order expansion around the SFT-midpoint, namely

$$\varphi_j = \varphi_{\frac{1}{2}} + \dot{\varphi}_{\frac{1}{2}} T_{\text{SFT}} \left( \frac{j}{N} - \frac{1}{2} \right) + \mathcal{O}(2). \quad (142)$$

Using this expansion, (141) now reads as

$$F_a^{X\alpha} \approx \hat{a} \Delta t \Delta f e^{-i2\pi\lambda} \sum_{k=0}^{N-1} \tilde{y}_k \sum_{j=0}^{N-1} e^{-i2\pi\kappa(k)j/N}, \quad (143)$$

where we defined

$$\begin{aligned} \lambda &\equiv \varphi_{\frac{1}{2}} - \frac{1}{2} \dot{\varphi}_{\frac{1}{2}} T_{\text{SFT}}, \\ \kappa(k) &\equiv \dot{\varphi}_{\frac{1}{2}} T_{\text{SFT}} - k. \end{aligned} \quad (144)$$

The last sum in (143) is simply a geometrical series, and so we find

$$\begin{aligned} \sum_{j=0}^{N-1} e^{-i2\pi\kappa j/N} &= \frac{1 - e^{-i2\pi\kappa}}{1 - e^{-i2\pi\kappa/N}} \\ &\stackrel{N \gg 1}{\approx} \frac{N}{2\pi} \left( \frac{\sin 2\pi\kappa}{\kappa} + i \frac{\cos 2\pi\kappa - 1}{\kappa} \right) \\ &\equiv \frac{N}{2\pi} P(\kappa(k)) = \frac{N}{2\pi} P_k. \end{aligned} \quad (145)$$

The function  $P(\kappa)$  is sometimes called ‘‘Dirichlet kernel’’, and it has the property of being strongly peaked around  $\kappa = 0$ , and so we can truncate the sum over  $k$  in (143) to a few terms  $\Delta k$  (referred to as **Dterms** in the code) on either side of  $k^*$ , corresponding to the frequency bin closest to the maximum of  $P(\kappa)$ , i.e. the bin closest to the solution of  $\kappa(k) = 0$ , namely

$$k^* \equiv \text{round} \left[ \dot{\varphi}_{\frac{1}{2}} T_{\text{SFT}} \right] = \text{round} \left[ \hat{f}(t_{\frac{1}{2}}) / \Delta f \right], \quad (146)$$

where  $\hat{f}(t)$  is the ‘‘effective’’ signal-frequency in the detector frame at time  $t$  (the time-derivative  $\dot{\varphi}$  refers to the time in the detector-frame!), which shows that generally we’ll have  $k^* \gg 1$ . With this approximation we finally find

$$F_a^{X\alpha} \approx \frac{1}{2\pi} \hat{a} e^{-i2\pi\lambda} \sum_{k=k^*-\Delta k}^{k^*+\Delta k} \tilde{y}_k P_k. \quad (147)$$

We’ll also need explicit expressions for  $\varphi_{\frac{1}{2}}$  and  $\dot{\varphi}_{\frac{1}{2}}$  in order to compute  $\lambda$  and  $\kappa(k)$ , defined in (144). For this we need the timing-function  $\tau(t)$ , which translates detector arrival times  $t$  to the SSB  $\tau$ . In the purely Newtonian approximation this would be given by (23), but in general the code uses a

full ephemeris-based relativistic timing model  $\tau(t)$  (in `LALBarycenter()`). Given this function, we define

$$\Delta\tau_{\frac{1}{2}} \equiv \tau(t_{\frac{1}{2}}) - \tau_{\text{ref}}, \quad (148)$$

$$\dot{\tau}_{\frac{1}{2}} \equiv \left. \frac{d\tau}{dt} \right|_{t_{\frac{1}{2}}} \quad (\approx 1 + \vec{v}_{\frac{1}{2}} \cdot \vec{n}/c), \quad (149)$$

(which are referred to as `(Multi)SSBtimes` in the code), and so the (full) phase-model (22) yields

$$\varphi_{\frac{1}{2}} = \sum_s \frac{f^{(s)}}{(s+1)!} \Delta\tau_{\frac{1}{2}}^{s+1}, \quad (150)$$

$$\dot{\varphi}_{\frac{1}{2}} = \dot{\tau}_{\frac{1}{2}} \sum_s \frac{f^{(s)}}{s!} \Delta\tau_{\frac{1}{2}}^s. \quad (151)$$

### 4.3 Efficient computation of the “atoms” $F_{\{a,b\}}^{X\alpha}$

The computation of (147) will be the most time-consuming part in this code, in particular the “hot loop” which is the sum over  $k$ . It is therefore important to compute these sums in the most efficient way possible.

First it will be convenient to relabel this sum using  $l(k) \equiv k - k_0$  with  $k_0 \equiv k^* - \Delta k$  being the leftmost bin in the sum, and so we write

$$\kappa_l \equiv \kappa(k(l)) = \kappa_0 - l, \quad (152)$$

where

$$\kappa_0 \equiv \text{rem} \left( \dot{\varphi}_{\frac{1}{2}} T_{\text{SFT}} \right) + \Delta k, \quad (153)$$

and where we defined the “remainder”

$$\text{rem}(x) \equiv x - \text{round}[x]. \quad (154)$$

Next we note that

$$\sin 2\pi\kappa_l = \sin 2\pi\kappa_0 \equiv s \quad (155)$$

$$\cos 2\pi\kappa_l - 1 = \cos 2\pi\kappa_0 - 1 \equiv c, \quad (156)$$

and so the Dirichlet-kernel (145) has the form

$$P_{k(l)} = \frac{s}{\kappa_0 - l} + i \frac{c}{\kappa_0 - l}. \quad (157)$$

Now let us look at the “hot loop” in (147), which we can express as

$$\chi \equiv \sum_{k=k_0}^{k_0+\mathcal{N}} \tilde{y}_k P_k = [sU - cV] + i [cU + sV], \quad (158)$$

where  $\mathcal{N} \equiv 2\Delta k - 1$ , and the two sums we need to evaluate are

$$U \equiv \sum_{l=0}^{\mathcal{N}} \frac{u_l}{p_l}, \quad V \equiv \sum_{l=0}^{\mathcal{N}} \frac{v_l}{p_l}, \quad (159)$$

with the further definitions

$$p_l \equiv \kappa_0 - l, \quad (160)$$

$$u_l \equiv \Re(\tilde{y}_{k_0+l}), \quad (161)$$

$$v_l \equiv \Im(\tilde{y}_{k_0+l}). \quad (162)$$

The above sums (159) are numerically not efficient as they consist of many divisions, which are slower than multiplications. This can be remedied with a clever algorithm suggested by Fekete Ākos: bringing the sums on a common denominator  $q_{\mathcal{N}}$ , we get

$$U = \frac{S_{\mathcal{N}}}{q_{\mathcal{N}}}, \quad V = \frac{T_{\mathcal{N}}}{q_{\mathcal{N}}}, \quad (163)$$

where

$$S_{\mathcal{N}} = u_0 p_1 p_2 \dots p_{\mathcal{N}} + p_0 u_1 p_1 \dots p_{\mathcal{N}} + \dots + p_0 p_1 \dots p_{\mathcal{N}-1} u_{\mathcal{N}}, \quad (164)$$

$$T_{\mathcal{N}} = v_0 p_1 p_2 \dots p_{\mathcal{N}} + p_0 v_1 p_1 \dots p_{\mathcal{N}} + \dots + p_0 p_1 \dots p_{\mathcal{N}-1} v_{\mathcal{N}}, \quad (165)$$

$$q_{\mathcal{N}} = p_0 p_1 \dots p_{\mathcal{N}}, \quad (166)$$

reducing the  $2\mathcal{N} + 2$  divisions to only 2. The required three components  $S_{\mathcal{N}}$ ,  $T_{\mathcal{N}}$  and  $q_{\mathcal{N}}$  can be computed efficiently using the following recurrence:

$$S_n = p_n S_{n-1} + q_{n-1} u_n, \quad (167)$$

$$T_n = p_n T_{n-1} + q_{n-1} v_n, \quad (168)$$

$$p_n = p_{n-1} - 1, \quad (169)$$

$$q_n = p_n q_{n-1}, \quad (170)$$

and the starting conditions

$$S_0 = u_0, \quad (171)$$

$$T_0 = v_0, \quad (172)$$

$$p_0 = \kappa_0, \quad (173)$$

$$q_0 = p_0. \quad (174)$$



The number of floating-point operations per iteration is 8, so in total we need  $8\mathcal{N} + 8$  operations (not counting one sin/cos), of which only 2 are divisions. In the previous “LALDemod” algorithm (e.g `ComputeFstat.c:1.19`)  $\chi$  was computed more directly resulting in  $12\mathcal{N}$  floating point operations, of which  $\mathcal{N}$  are divisions!

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